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RANDOM COVERS OF FINITE HOMOGENEOUS LATTICES

We develop and extend some results for the scheme of independent random elements distributed on a finite lattice. In particular, we introduce the concept of the cover of a homogeneous lattice L_n of rank n and derive the exact equations and estimations for the number of covers with a given number of blocks and for the total covers number of the lattice L_n . A theorem about the asymptotic normality of the blocks number in a random equiprobable cover of the lattice L_n is proved. The concept of the cover index of the lattice L_n , that extend the notion of the cover index of a finite set by its independent random subsets, is introduced. Applying the lattice moments method, the limit distribution as $n \rightarrow \infty$ for the cover index of a subspace lattice of the n -dimensional vector space over a finite field is determined.

1. BASIC NOTIONS AND PRELIMINARY RESULTS

In this paper, we use notions and results from [1, 2]. We refer also to [3, 4] for the terminology and detailed exposition of finite lattices theory.

Let $L = \{L_n : n = 0, 1, \dots\}$ be a sequence of finite lattices. Denote the Moebius function, the maximal and minimal elements of the lattice L_n by $\mu_n, 1_n$, and 0_n respectively. We assume that the sequence L satisfies the following homogeneity conditions (see [4]):

- (a) L_n is a graduate lattice with rank function r , where $r(1_n) = n$ for any $n = 0, 1, \dots$;
- (b) for any $X \in L_n$ such that $r(X) = n - k, k \in \overline{0, n}$, the interval $[X, 1_n]$ is isomorphic to the lattice L_k .

Let

$$(1) \quad w(n, k) = \sum_{a \in L_n: r(a)=k} \mu_n(0_n, a), \quad W(n, k) = \sum_{a \in L_n: r(a)=k} 1,$$

where $k \in \overline{0, n}, n = 0, 1, \dots$; $w(n, k) = W(n, k) = 0$ otherwise. The numbers $w(n, k)$ and $W(n, k)$ are called the k -th level numbers of the lattice L_n of the first kind and of the second kind respectively [3, 4].

By $\chi_n(z)$, we denote the characteristic polynomial of the lattice L_n ,

$$(2) \quad \chi_n(z) = \sum_{k=0}^n w(n, n-k) z^k, \quad n = 0, 1, \dots$$

In what follows, we assume that there exists a sequence of real numbers $a_i \geq 1$ ($i = 1, 2, \dots$) such that

$$(3) \quad \chi_n(z) = \prod_{i=1}^n (z - a_i), \quad n = 1, 2, \dots, \quad \chi_0(z) \equiv 1.$$

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Note that condition (3) holds for the characteristic polynomial of any finite super-solvable geometric lattice (see [4, 5]).

We consider the following most important examples of homogeneous lattices which satisfy condition (3). Other examples can be found in [3, 4, 5].

1. $L_n = B(n)$, where $B(n)$ is the set of all subsets of $N_n = \{1, 2, \dots, n\}$. The rank function, level numbers, and characteristic polynomial of the lattice $B(n)$ are equal, respectively, to $r(X) = \#X$ (where $\#X$ denotes the cardinality of a set $X \in B(n)$),

$$w(n, k) = (-1)^k \binom{n}{k}, \quad W(n, k) = \binom{n}{k}, \quad \chi_n(z) = (z-1)^n, \quad k \in \overline{0, n}, \quad n = 0, 1, \dots$$

2. $L_n = L(n, q)$ is a subspace lattice of the n -dimensional vector space $V(n, q)$ over a field with q elements. In this case, the rank $r(X)$ of a subspace $X \in L_n$ is equal to the dimension of X ,

$$w(n, k) = (-1)^k q^{\binom{k}{2}} \begin{bmatrix} n \\ k \end{bmatrix}_q, \quad W(n, k) = \begin{bmatrix} n \\ k \end{bmatrix}_q, \quad \chi_n(z) = \prod_{i=1}^n (z - q^{i-1}),$$

where

$$\begin{bmatrix} n \\ k \end{bmatrix}_q = \frac{(q^{-1})_n}{(q^{-1})_k (q^{-1})_{n-k}} q^{k(n-k)}$$

is the Gauss coefficient (the number of k -dimensional subspaces of the vector space $V(n, q)$, $k \in \overline{0, n}$), $(q^{-1})_n = (1 - q^{-1})(1 - q^{-2}) \dots (1 - q^{-n})$, $n = 0, 1, \dots$.

3. $L_n = \mathfrak{R}(n+1)$ is the lattice of all partitions of the set $N_{n+1} = \{1, 2, \dots, n+1\}$. The rank of a partition $\pi \in \mathfrak{R}(n+1)$ is equal to $r(\pi) = n+1 - b(\pi)$, where $b(\pi)$ is the number of blocks in π . The level numbers and the characteristic polynomial of the lattice $\mathfrak{R}(n+1)$ are, respectively, equal to

$$w(n, k) = s(n+1, n+1-k), \quad W(n, k) = S(n+1, n+1-k), \quad \chi_n(z) = \prod_{i=1}^n (z-i),$$

where $k \in \overline{0, n}$, $n = 0, 1, \dots$, $s(\cdot, \cdot)$ and $S(\cdot, \cdot)$ are Stirling numbers of the first kind and of the second kind, respectively.

Let's consider a sequence of random variables ξ_0, ξ_1, \dots , where ξ_n takes values in the set $\{0, 1, \dots, n\}$, $n = 0, 1, \dots$. Let $p_{n,k} = \mathbf{P}\{\xi_n = k\}$, $B_{n,k} = \mathbf{E}W(\xi_n, \xi_n - k)$, $k \in \overline{0, n}$, $n = 0, 1, \dots$. We call $B_{n,k}$ the k -th lattice moment of the random variable ξ_n [2].

The following statement was proved in [2].

Statement 1. *Let L be a sequence of finite homogeneous lattices with the level numbers (1) and the characteristic polynomials (2) that satisfies condition (3). Then*

1. *We have the following mutually inverse relations:*

$$(4) \quad B_{n,l} = \sum_{k=l}^n p_{n,k} W(k, k-l), \quad p_{n,l} = \sum_{k=l}^n B_{n,k} w(k, k-l), \quad l \in \overline{0, n}, \quad n = 0, 1, \dots$$

2. *For the generating function of the random variable ξ_n , the following equalities hold:*

$$p_n(z) = \sum_{l=0}^n p_{n,l} z^l = \sum_{k=0}^n B_{n,k} \chi_k(z), \quad n = 0, 1, \dots,$$

$$(5) \quad \sum_{k=0}^{2\nu+1} B_{n,k} \chi_k(z) \leq p_n(z) \leq \sum_{k=0}^{2\nu} B_{n,k} \chi_k(z), \quad z \in [0, 1],$$

where $\nu = 0, 1, \dots$

Assuming that $z = 0$ in (5), we obtain the following inequalities:

$$\sum_{k=0}^{2\nu+1} B_{n,k} w(k, k) \leq p_{n,0} \leq \sum_{k=0}^{2\nu} B_{n,k} w(k, k), \quad \nu = 0, 1, \dots,$$

and, for $\nu = 0$,

$$(6) \quad 1 - B_{n,1} \leq p_{n,0} \leq 1.$$

2. RANDOM COVERS OF THE LATTICE L_n

By $\Lambda_{n,T}$, we denote the collection of all sets $X = \{X_1, \dots, X_T\}$ such that $X_1 \dots X_T$ are different non-zero elements of the lattice L_n . Let's assign the equiprobable distribution to the $\Lambda_{n,T}$, assuming that

$$(7) \quad p(X) = \binom{\lambda_n - 1}{T}^{-1}, \quad X = \{X_1, \dots, X_T\} \in \Lambda_{n,T},$$

where $\lambda_n = \#L_n$. Put

$$(8) \quad \lambda_n(Y) = \#[0_n, Y], \quad Y \in L_n,$$

$$(9) \quad \lambda^{(n-1)} = \max\{\lambda_n(Y) : Y \in L_n, r(Y) = n - 1\}, \quad n = 0, 1, \dots$$

Definition 1. A set $X = \{X_1, \dots, X_T\} \in \Lambda_{n,T}$ will be called a T -block cover of the lattice L_n (and its elements be called blocks of the cover X), if $X_1 \vee \dots \vee X_T = 1_n$.

By $r_{n,T} = n - r(X_1 \vee \dots \vee X_T)$, we denote the random variable equal to the co-rank of the join of X_1, \dots, X_T , where $X = \{X_1, \dots, X_T\}$ is a random element distributed according to (7). Let's denote

$$p_{n,l}^{(T)} = \mathbf{P}\{r_{n,T} = l\}, \quad B_{n,l}^{(T)} = \mathbf{E}W(r_{n,T}, r_{n,T} - l), \quad l \in \overline{0, n}, \quad n = 0, 1, \dots$$

We denote the T -block cover number and the total block cover number of the lattice L_n by $D_{n,T}$ and $D_n = \sum_{T=1}^{\lambda_n-1} D_{n,T}$, respectively.

For the case $L_n = B(n)$, exact formulas and estimations of the numbers $D_{n,T}$, D_n are obtained in [6, p. 269]. The following statement extends these results.

Statement 2. *The following relations hold:*

$$(10) \quad D_{n,T} = \binom{\lambda_n - 1}{T} \sum_{k=0}^n B_{n,k}^{(T)} w(k, k), \quad D_n = \sum_{k=0}^n w(k, k) \sum_{T=1}^{\lambda_n-1} \binom{\lambda_n - 1}{T} B_{n,k}^{(T)},$$

$$(11) \quad \binom{\lambda_n - 1}{T} - B_{n,1}^{(T)} \binom{\lambda_n - 1}{T} \leq D_{n,T} \leq \binom{\lambda_n - 1}{T},$$

$$(12) \quad 2^{\lambda_n-1} - 1 - \sum_{T=1}^{\lambda_n-1} \binom{\lambda_n - 1}{T} B_{n,1}^{(T)} \leq D_{n,T} \leq 2^{\lambda_n-1} - 1.$$

As this takes place, the following equation holds for any $l \in \overline{0, n}$, $T \in \overline{1, \lambda_n - 1}$:

$$(13) \quad B_{n,l}^{(T)} = \binom{\lambda_n - 1}{T}^{-1} \sum_{\substack{Y \in L_n: \\ r(Y) = n-l}} \binom{\lambda_n(Y) - 1}{T}, \quad l \in \overline{0, n}, \quad T \in \overline{1, \lambda_n - 1}.$$

Proof. The first equality in (10) is immediate from (4) and the equality

$$D_{n,T} = \binom{\lambda_n - 1}{T} p_{n,0}^{(T)};$$

the second equality in (10) follows from the first one. Inequalities (11) follow from (6) and the first equation in (10); inequalities (12) are obtained by the summation of (11) over $T \in \overline{1, \lambda_n - 1}$. Finally, the proof of (13) is similar to the proof of Theorem 1 [1] applying the Moebius inversion formula.

3. ASYMPTOTIC BEHAVIOR OF THE BLOCK NUMBER DISTRIBUTION IN A RANDOM EQUIPROBABLE COVER OF THE LATTICE L_n

Let ζ_n denote the random variable equal to the block number in a random equiprobable cover of the lattice L_n . Let's prove the following theorem generalizing a result from [7].

Theorem 1. *Suppose that*

$$(14) \quad \overline{\lim}_{n \rightarrow \infty} \lambda^{(n-1)} (\lambda_n)^{-1} < 1,$$

where $\lambda^{(n-1)}$ is defined by (9). Then

$$(15) \quad D_n = 2^{\lambda_n - 1} (1 + o(1)), \quad n \rightarrow \infty,$$

$$(16) \quad \mathbf{P}\{\zeta_n = T\} = \frac{D_{n,T}}{D_n} = \frac{1}{\sqrt{\frac{\pi}{2}(\lambda_n - 1)}} \exp\left\{-\frac{(x_{n,T})^2}{2}\right\} (1 + o(1)),$$

provided n and T tend to infinity in such a way that

$$(17) \quad x_{n,T} \stackrel{\text{def}}{=} \frac{T - \frac{1}{2}(\lambda_n - 1)}{\frac{1}{2}\sqrt{\lambda_n - 1}} = o(\lambda_n^{\frac{1}{6}}).$$

Under condition (14), the remainder term on the right-hand side of (16) tends uniformly to zero for all T such that $x_{n,T}$ lies in any fixed finite interval.

Proof. First we show that, under assumption (14),

$$(18) \quad 2^{-(\lambda_n - 1)} \sum_{\substack{Y \in L_n: \\ r(Y) = n-1}} (2^{\lambda_n(Y) - 1} - 1) = o(1), \quad n \rightarrow \infty.$$

Applying the second equality from (1), we obtain

$$\lambda_n = \#L_n = \sum_{k=0}^n W(n, k) > W(n, n-1).$$

Whence and from (14), we have

$$(19) \quad \begin{aligned} 2^{-(\lambda_n - 1)} \sum_{\substack{Y \in L_n: \\ r(Y) = n-1}} (2^{\lambda_n(Y) - 1} - 1) &< \sum_{\substack{Y \in L_n: \\ r(Y) = n-1}} 2^{\lambda_n(Y)} \leq 2^{\lambda^{(n-1)} - \lambda_n} W(n, n-1) < \\ &< \lambda_n 2^{\lambda^{(n-1)} - \lambda_n} = 2^{-\lambda_n \left(1 - \frac{\lambda^{(n-1)}}{\lambda_n} - \frac{\log \lambda_n}{\lambda_n}\right)} = o(1), \quad n \rightarrow \infty. \end{aligned}$$

So equality (18) is proved.

Now, taking into account (12) and (13), we obtain

$$2^{\lambda_n-1} - 1 - \sum_{T=1}^{\lambda_n-1} \binom{\lambda_n-1}{T} B_{n,1}^{(T)} = 2^{\lambda_n-1} - \sum_{\substack{Y \in L_n: \\ r(Y)=n-1}} (2^{\lambda_n(Y)-1} - 1) \leq D_n \leq 2^{\lambda_n-1} - 1.$$

So, applying (18), we arrive at (15).

To prove equality (16), we set

$$\alpha(n, T) = \frac{1}{\sqrt{\frac{\pi}{2}(\lambda_n - 1)}} \exp \left\{ -\frac{(x_{n,T})^2}{2} \right\}.$$

Then

$$(20) \leq \left| \frac{D_{n,T}}{D_n} \alpha(n, T)^{-1} - 1 \right| \leq \left| \frac{D_{n,T}}{D_n} \alpha(n, T)^{-1} - \frac{1}{D_n} \binom{\lambda_n-1}{T} \alpha(n, T)^{-1} \right| + \left| \frac{1}{D_n} \binom{\lambda_n-1}{T} \alpha(n, T)^{-1} - 1 \right|.$$

From (11) and (13), we obtain that the augend on the right-hand side of (20) is not greater than

$$\alpha(n, T)^{-1} \frac{1}{D_n} \sum_{\substack{Y \in L_n: \\ r(Y)=n-1}} \binom{\lambda_n(Y)-1}{T} < \alpha(n, T)^{-1} \frac{1}{D_n} \lambda_n 2^{\lambda^{(n-1)}} =$$

$$\alpha(n, T)^{-1} \lambda_n 2^{\lambda^{(n-1)} - (\lambda_n - 1)} (1 + o(1)), \quad n \rightarrow \infty$$

[see relations (19) and (15)]. Further, applying (17) and (19), we obtain

$$\alpha(n, T)^{-1} \lambda_n 2^{\lambda^{(n-1)} - (\lambda_n - 1)} = O \left(\exp \left\{ -\frac{(x_{n,T})^2}{2} \right\} (\lambda_n)^{\frac{3}{2}} 2^{\lambda^{(n-1)} - \lambda_n} \right) = o(1), \quad n, T \rightarrow \infty.$$

Due to these relations, the augend on the right-hand side of (20) tends to zero as $n, T \rightarrow \infty$, and this convergence is uniform for all T , for which $x_{n,T}$ lies in any fixed finite interval.

To estimate the addend on the right-hand side of (20), we employ equality (15) and the Moivre – Laplace local theorem. Thus, we obtain that

$$\left| \frac{1}{D_n} \binom{\lambda_n-1}{T} \alpha(n, T)^{-1} - 1 \right| = o(1), \quad n, T \rightarrow \infty,$$

where $o(1)$ tends to zero uniformly for all T , for which $x_{n,T}$ lies in any fixed finite interval.

So equality (16) is completely proved, and so is the theorem.

Corollary. Under condition (14), the sequence of random variables $\{\zeta_n : n = 0, 1, \dots\}$ is asymptotically normal with parameters $\frac{1}{2}(\lambda_n - 1)$, $\frac{1}{2}\sqrt{\lambda_n - 1}$:

$$\mathbf{P} \left\{ \frac{\zeta_n - \frac{1}{2}(\lambda_n - 1)}{\frac{1}{2}\sqrt{\lambda_n - 1}} \leq x \right\} \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{t^2}{2}} dt, \quad n \rightarrow \infty.$$

Notice that condition (14) is fulfilled if L_n is one of the lattices $B(n)$, $L(n, q)$, $\mathfrak{R}(n+1)$ from Section 1. For $B(n)$, Theorem 1 and its Corollary were earlier obtained by Sachkov [7]. It is evident that, in this case, the limit on the left-hand side of (14) equals $1/2$.

If $L_n = L(n, q)$, $n = 0, 1, \dots$, then $\lambda_n = G_n$, $\lambda^{(n-1)} = G_{n-1}$, where $G_n = \sum_{k=0}^n \begin{bmatrix} n \\ k \end{bmatrix}_q$ is the total subspace number of the vector space $V(n, q)$ (the Galois number) [6]. In this case, inequality (14) follows from the asymptotic formula [8]

$$(21) \quad G_n - 1 = \theta_{n(2)}(q)q^{\frac{n^2}{4}}(1 + O(q^{-\frac{n}{2}})), \quad n \rightarrow \infty,$$

where $n(2)$ denotes the residue n modulo 2,

$$(22) \quad \theta_0(q) = \frac{1}{(q^{-1})_\infty} \sum_{n=-\infty}^{\infty} q^{-n^2}, \quad \theta_1(q) = \frac{1}{(q^{-1})_\infty} \sum_{n=-\infty}^{\infty} q^{-(n-\frac{1}{2})^2},$$

$$(q^{-1})_\infty = \prod_{m=1}^{\infty} (1 - q^{-m}).$$

In the case of $L_n = \mathfrak{R}(n+1)$, $n = 0, 1, \dots$, (14) follows from the asymptotic formula for Bell numbers (see, for example, [6, p. 297]).

4. PROBABILITY DISTRIBUTION ASYMPTOTIC BEHAVIOR OF THE COVER INDEX OF A FINITE HOMOGENEOUS LATTICE

Let $\Xi = X_1, X_2, \dots$ be a sequence of independent random elements of the lattice L_n .

Definition 2. The cover index of the lattice L_n by elements of the sequence Ξ is the least $\theta = \theta(n, \Xi) \in N$ such that

$$(23) \quad X_1 \vee \dots \vee X_\theta = 1_n.$$

This definition extends the concept of the n -set cover index by its independent random subsets (see [9]).

Taking into account the equality $\{\theta(n, \Xi) \leq T\} = \{r_{n,T} = 0\}$, where $r_{n,T} = n - r(X_1 \vee \dots \vee X_T)$ is the co-rank of the join of T first elements of the sequence Ξ , we can to apply the lattice moments method [2] for obtaining the limit distribution of the random sequence $\theta(n, \Xi)$.

For any natural n, T and $l \in \overline{0, n}$, we set

$$(24) \quad B_{n,l}^{(T)} = \sum_{\substack{Y \in L_n: \\ r(Y) = n-l}} \mathbf{P}\{X_1 \leq Y\} \dots \mathbf{P}\{X_T \leq Y\}.$$

Theorem 2. Let n and $T = T(n)$ be such that, for any $l = 0, 1, \dots$,

$$(25) \quad B_{n,l}^{(T)} \rightarrow B_l < \infty, \quad n \rightarrow \infty.$$

Then, if the series

$$(26) \quad B(z) \stackrel{\text{def}}{=} \sum_{l=0}^{\infty} B_l \chi(z) = \sum_{l=0}^{\infty} B_l \left(\sum_{k=0}^l w(l, l-k) z^k \right)$$

is uniformly convergent in a neighborhood of zero, then

$$(27) \quad \lim_{n \rightarrow \infty} \mathbf{P}\{\theta(n, \Xi) \leq T\} = \sum_{l=0}^{\infty} w(l, l) B_l.$$

Proof. It follows from Theorem 2 in [1] that numbers (24) are equal to the lattice moments of the random variable $r_{n,T}$. Thus, expression (27) follows directly from equalities (25) and $\{\theta(n, \Xi) \leq T\} = \{r_{n,T} = 0\}$ and the statement of Theorem 2 in [2].

Theorem 2 yields the expression for the limit distribution law (as $n \rightarrow \infty$) of the subspace lattice cover index of the n -dimensional vector space over a field with q elements.

Theorem 3. *Let Ξ be a sequence of independent and equiprobable random subspaces in the vector space $V(n, q)$. Then*

$$(28) \quad \lim_{\substack{n \rightarrow \infty \\ n \equiv i \pmod{2}}} \mathbf{P}\{\theta(n, \Xi) \leq 2\} = \frac{1}{2} \left(\frac{1 + C_i}{(-q^{-1/2})_\infty} + \frac{1 - C_i}{(q^{-1/2})_\infty} \right),$$

where

$$(29) \quad i \in \{0, 1\}, C_0 = 4q^{-1/2}, C_1 = 1/4q^{1/2}, (x)_\infty = \prod_{m=0}^{\infty} (1 - xq^{-m}).$$

In addition,

$$(30) \quad \lim_{n \rightarrow \infty} \mathbf{P}\{\theta(n, \Xi) \leq 3\} = 1.$$

Proof. From (24), we get

$$(31) \quad B_{n,l}^{(T)} = \begin{bmatrix} n \\ l \end{bmatrix}_q \left(\frac{G_{n-l}}{G_n} \right)^T, \quad n, T \in \{1, 2, \dots\}, l \in \overline{0, n}.$$

Hence, applying (21), it is easy to obtain that, for $T = 3$,

$$B_0^{(3)} \stackrel{\text{def}}{=} \lim_{n \rightarrow \infty} B_{n,0}^{(3)} = 1, B_l^{(3)} = \lim_{n \rightarrow \infty} B_{n,l}^{(3)} = 0, \quad l = 1, 2, \dots$$

Note that series (26) corresponding to values $B_l = B_l^{(3)}$, $l = 0, 1, \dots$ is uniformly convergent over the complex plane. Hence, we obtain (30) from (27).

For $T = 2$, applying (31) and (21), we get

$$B_{n,l}^{(2)} = \frac{q^{-l^2/2}}{(q^{-1})_l} \left(\frac{\theta_{(n-l)(2)}}{\theta_{n(2)}} \right)^2 (1 + o(1)), \quad n \rightarrow \infty, l = 0, 1, \dots,$$

Hence, taking into account the equality $\theta_1(q) = 2q^{-1/4}\theta_0(q)$ emerging from the Jacobi identity (see, for example, [10]), we obtain

$$(32) \quad B_{0,l} \stackrel{\text{def}}{=} \lim_{\substack{n \rightarrow \infty \\ n \equiv 0 \pmod{2}}} B_{n,l}^{(2)} = \frac{q^{-l^2/2}}{(q^{-1})_l}, \quad \text{if } l \equiv 0 \pmod{2},$$

$$B_{0,l} = 4q^{-1/2} \frac{q^{-l^2/2}}{(q^{-1})_l}, \quad \text{if } l \equiv 1 \pmod{2},$$

$$B_{1,l} \stackrel{\text{def}}{=} \lim_{\substack{n \rightarrow \infty \\ n \equiv 1 \pmod{2}}} B_{n,l}^{(2)} = \frac{q^{-l^2/2}}{(q^{-1})_l}, \quad \text{if } l \equiv 0 \pmod{2},$$

$$(33) \quad B_{1,l} = 1/4q^{1/2} \frac{q^{-l^2/2}}{(q^{-1})_l}, \quad \text{if } l \equiv 1 \pmod{2};$$

As far as series (26) corresponding to each sequence $\{B_l = B_{i,l} : l = 0, 1, \dots\}$, $i \in \{0, 1\}$, is uniformly convergent over the whole complex plane, we obtain, by applying (27), the following equalities:

$$(34) \quad \lim_{\substack{n \rightarrow \infty \\ n \equiv i \pmod{2}}} \mathbf{P}\{\theta(n, \Xi) \leq 2\} = \sum_{l=0}^{\infty} w(l, l) B_{i,l} = \sum_{l=0}^{\infty} (-1)^l q^{\binom{l}{2}} B_{i,l}, \quad i \in \{0, 1\}.$$

Substituting (32), (33), respectively, in (34) and applying the equalities

$$\sum_{\substack{k \geq 0: \\ k \equiv 0 \pmod{2}}} (-1)^k \frac{q^{-k/2}}{(q^{-1})_k} = \frac{1}{2} \left(\frac{1}{(-q^{-1/2})_\infty} + \frac{1}{(q^{-1/2})_\infty} \right),$$

$$\sum_{\substack{k \geq 0: \\ k \equiv 1 \pmod{2}}} (-1)^k \frac{q^{-k/2}}{(q^{-1})_k} = \frac{1}{2} \left(\frac{1}{(-q^{-1/2})_\infty} - \frac{1}{(q^{-1/2})_\infty} \right)$$

following from the identity $\sum_{k=0}^{\infty} \frac{t^k}{(q^{-1})_k} = \frac{1}{(t)_\infty}$, $|t| < 1$ (see [10]), we obtain equality (28) after simple transformations.

So, the theorem is proved.

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